NPFL129, Lecture 3



Perceptron and Logistic Regression

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unless otherwise stated



After this lecture you should be able to

- Think about binary classification using **geometric intuition** and use the **perceptron algorithm**.
- Define the **main concepts of information theory** (entropy, cross-entropy, KL-divergence) and prove their properties.
- Derive training objectives using the **maximum likelihood principle**.
- Implement and use **logistic regression** for binary classification with SGD.

Binary Classification



Binary classification is a classification in two classes.

The simplest way to evaluate classification is **accuracy**, which is the ratio of input examples that were classified correctly – i.e., where the predicted class and the target class match.

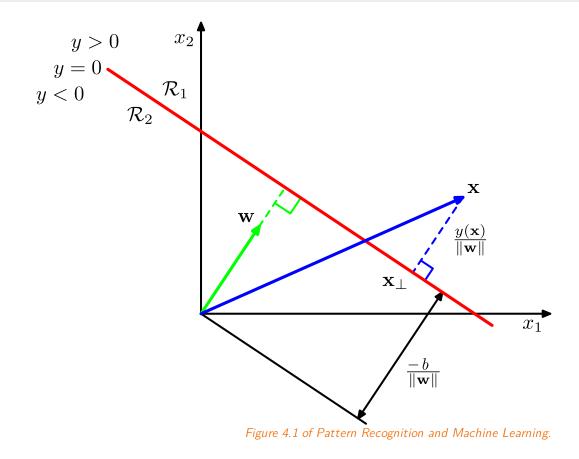
To extend linear regression to binary classification, we might seek a **threshold** and then classify an input as negative/positive depending on whether $y(\boldsymbol{x}; \boldsymbol{w}) = \boldsymbol{x}^T \boldsymbol{w} + b$ is smaller/larger than a given threshold.

Zero value is usually used as the threshold, both because of symmetry and also because the **bias** parameter acts as a trainable threshold anyway.

The set of points with prediction 0 is called a **decision boundary**.

Geometric Intuition





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Perceptron

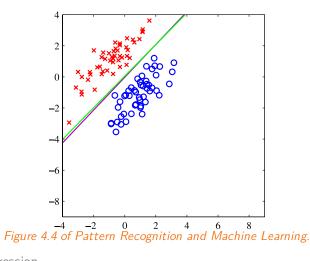
The perceptron algorithm is probably the oldest one for training weights of a binary classification. Assuming the target value $t \in \{-1, +1\}$, the goal is to find weights w such that for all train data,

$$ext{sign}(y(oldsymbol{x}_i;oldsymbol{w})) = ext{sign}(oldsymbol{x}_i^Toldsymbol{w}) = t_i,$$

or equivalently,

$$t_i y(oldsymbol{x}_i;oldsymbol{w}) = t_i oldsymbol{x}_i^T oldsymbol{w} > 0.$$

Note that a set is called **linearly separable**, if there exists a weight vector \boldsymbol{w} such that the above equation holds.



InformationTheory

Perceptron



The perceptron algorithm was invented by Rosenblatt in 1958.

Input: Linearly separable dataset ($m{X} \in \mathbb{R}^{N \times D}$, $m{t} \in \{-1, +1\}^N$). Output: Weights $m{w} \in \mathbb{R}^D$ such that $t_i m{x}_i^T m{w} > 0$ for all i.

- $oldsymbol{w} \leftarrow oldsymbol{0}$
- until all examples are classified correctly, process example i:
 - $\circ y \leftarrow oldsymbol{x}_i^T oldsymbol{w}$
 - \circ if $t_iy \leq 0$ (incorrectly classified example):
 - $\boldsymbol{w} \leftarrow \boldsymbol{w} + t_i \boldsymbol{x}_i$

We will prove that the algorithm always arrives at some correct set of weights \boldsymbol{w} if the training set is linearly separable.

Proof of Perceptron Convergence

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Perceptron

Probability Basics

Let w_* be some weights correctly classifying (separating) the training data, and let w_k be the weights after k nontrivial updates of the perceptron algorithm, with w_0 being 0.

We will prove that the angle lpha between $oldsymbol{w}_*$ and $oldsymbol{w}_k$ decreases at each step. Note that

$$\cos(lpha) = rac{oldsymbol{w}_*^Toldsymbol{w}_k}{\|oldsymbol{w}_*\|\cdot\|oldsymbol{w}_k\|}.$$





Proof of Perceptron Convergence

Assume that the maximum norm of any training example $||\boldsymbol{x}||$ is bounded by R, and that γ is the minimum margin of \boldsymbol{w}_* , so for each training example (\boldsymbol{x}, t) , $t\boldsymbol{x}^T\boldsymbol{w}_* \geq \gamma$. First consider the dot product of \boldsymbol{w}_* and \boldsymbol{w}_k :

$$oldsymbol{w}_*^Toldsymbol{w}_k = oldsymbol{w}_*^T(oldsymbol{w}_{k-1} + t_koldsymbol{x}_k) \geq oldsymbol{w}_*^Toldsymbol{w}_{k-1} + \gamma.$$

By iteratively applying this equation, we get

$$oldsymbol{w}_*^Toldsymbol{w}_k\geq k\gamma.$$

Now consider the length of \boldsymbol{w}_k :

$$\|m{w}_k\|^2 = \|m{w}_{k-1} + t_km{x}_k\|^2 = \|m{w}_{k-1}\|^2 + 2t_km{x}_k^Tm{w}_{k-1} + \|m{x}_k\|^2.$$

Because \boldsymbol{x}_k was misclassified, we know that $t_k \boldsymbol{x}_k^T \boldsymbol{w}_{k-1} \leq 0$, so $\|\boldsymbol{w}_k\|^2 \leq \|\boldsymbol{w}_{k-1}\|^2 + R^2$. When applied iteratively, we get $\|\boldsymbol{w}_k\|^2 \leq k \cdot R^2$.

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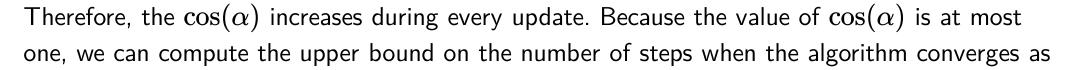


Proof of Perceptron Convergence



Putting everything together, we get

$$\cos(lpha) = rac{oldsymbol{w}_*^Toldsymbol{w}_k}{\|oldsymbol{w}_*\|\cdot\|oldsymbol{w}_k\|} \geq rac{k\gamma}{\sqrt{kR^2}\|oldsymbol{w}_*\|}.$$



$$1 \geq rac{\sqrt{k} \gamma}{\sqrt{R^2} \|oldsymbol{w}_*\|} ext{ or } k \leq rac{R^2 \|oldsymbol{w}_*\|^2}{\gamma^2}$$

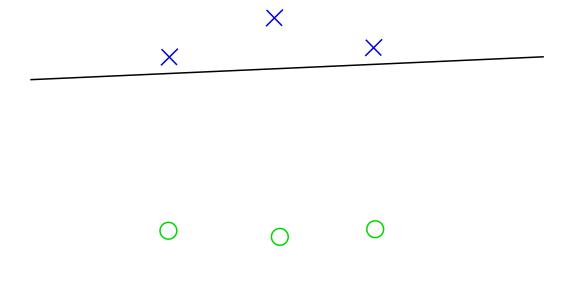
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Perceptron Issues



Perceptron has several drawbacks:

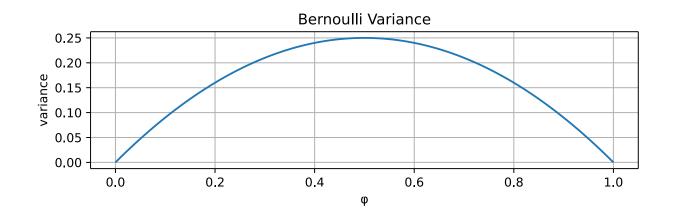
- If the input set is not linearly separable, the algorithm never finishes.
- The algorithm performs only prediction, it is not able to return the probabilities of predictions.
- Most importantly, Perceptron algorithm finds *some* solution, not necessarily a good one, because once it finds some, it cannot perform any more updates.



Bernoulli Distribution

The Bernoulli distribution is a distribution over a binary random variable. It has a single parameter $\varphi \in [0, 1]$, which specifies the probability that the random variable is equal to 1.

$$egin{aligned} P(x) &= arphi^x (1-arphi)^{1-x} \ &\mathbb{E}[x] &= arphi \ &\mathrm{Var}(x) &= arphi(1-arphi) \end{aligned}$$



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Common Probability Distributions

Categorical Distribution

Extension of the Bernoulli distribution to random variables taking one of K different discrete outcomes. It is parametrized by $p \in [0, 1]^K$ such that $\sum_{i=0}^{K-1} p_i = 1$.

We represent outcomes as vectors $\in \{0,1\}^K$ in the **one-hot encoding**. Therefore, an outcome $x \in \{0, 1, \dots, K-1\}$ is represented as a vector

$$\mathbf{1}_x \stackrel{ ext{def}}{=} ig([i=x]ig)_{i=0}^{K-1} = ig(\underbrace{0,\ldots,0}_x,1,\underbrace{0,\ldots,0}_{K-x-1}ig).$$

The outcome probability, mean, and variance are very similar to the Bernoulli distribution.

$$egin{aligned} P(oldsymbol{x}) &= \prod_{i=0}^{K-1} p_i^{x_i} \ \mathbb{E}[x_i] &= p_i \ \mathrm{Var}(x_i) &= p_i(1-p_i) \end{aligned}$$

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Self Information

Amount of **surprise** when a random variable is sampled.

- Should be zero for events with probability 1.
- Less likely events are more surprising.
- Independent events should have **additive** information.

$$I(x) \stackrel{\scriptscriptstyle{ ext{def}}}{=} -\log P(x) = \log rac{1}{P(x)}$$



Information Theory

Entropy

Amount of **surprise** in the whole distribution.

$$H(P) \stackrel{ ext{def}}{=} \mathbb{E}_{\mathrm{x} \sim P}[I(x)] = -\mathbb{E}_{\mathrm{x} \sim P}[\log P(x)]$$

- for discrete P: $H(P) = -\sum_x P(x) \log P(x)$
- for continuous P: $H(P) = -\int P(x) \log P(x) \, \mathrm{d}x$

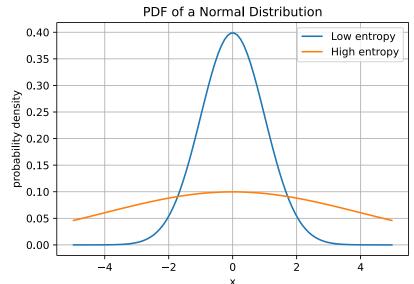
Because $\lim_{x o 0} x \log x = 0$, for P(x) = 0 we consider $P(x) \log P(x)$ to be zero.

Note that in the continuous case, the continuous entropy (also called *differential entropy*) has slightly different semantics, for example, it can be negative.

For binary logarithms, the entropy is measured in **bits**. × However, from now on, all logarithms are *natural logarithms* with base *e* (and then the entropy is measured in units called **nats**).

MLE

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Cross-Entropy

$$H(P,Q) \stackrel{ ext{def}}{=} -\mathbb{E}_{\mathrm{x}\sim P}[\log Q(x)]$$

Gibbs Inequality

- $H(P,Q) \ge H(P)$
- $H(P) = H(P,Q) \Leftrightarrow P = Q$

Proof: Consider $H(P) - H(P,Q) = \sum_{x} P(x) \log \frac{Q(x)}{P(x)}$.

Using the fact that $\log x \leq (x-1)$ with equality only for x=1, we get

$$\sum_x P(x)\lograc{Q(x)}{P(x)}\leq \sum_x P(x)\left(rac{Q(x)}{P(x)}-1
ight)=\sum_x Q(x)-\sum_x P(x)=0.$$

For the equality to hold, $\frac{Q(x)}{P(x)}$ must be 1 for all x, i.e., P = Q.

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Information Theory



Kullback-Leibler Divergence (KL Divergence)

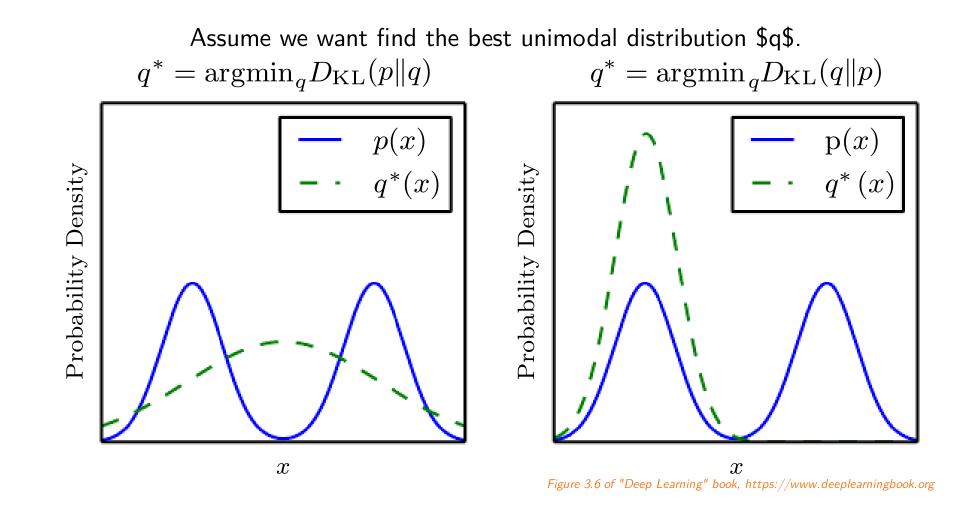
Sometimes also called **relative entropy**.

$$D_{ ext{KL}}(P\|Q) \stackrel{ ext{def}}{=} H(P,Q) - H(P) = \mathbb{E}_{ ext{x} \sim P}[\log P(x) - \log Q(x)]$$

- consequence of Gibbs inequality: $D_{ ext{KL}}(P\|Q) \geq 0$, $D_{ ext{KL}}(P\|Q) = 0$ iff P = Q
- generally $D_{ ext{KL}}(P\|Q)
 eq D_{ ext{KL}}(Q\|P)$

Nonsymmetry of KL Divergence





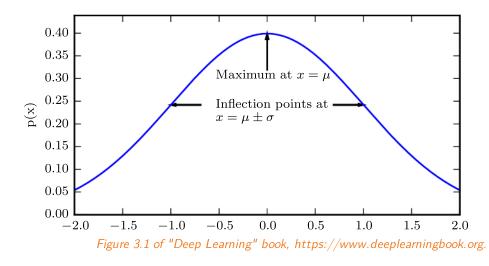
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Normal (or Gaussian) Distribution

Distribution over real numbers, parametrized by a mean μ and variance σ^2 :

$$\mathcal{N}(x;\mu,\sigma^2) = \sqrt{rac{1}{2\pi\sigma^2}} \exp\left(-rac{(x-\mu)^2}{2\sigma^2}
ight)\,.$$

For standard values $\mu=0$ and $\sigma^2=1$ we get $\mathcal{N}(x;0,1)=\sqrt{rac{1}{2\pi}}e^{-rac{x^2}{2}}$.



Why Normal Distribution



Central Limit Theorem

The sum of independent identically distributed random variables with finite variance converges to normal distribution.

Principle of Maximum Entropy

Given a set of constraints, a distribution with maximal entropy fulfilling the constraints can be considered the most general one, containing as little additional assumptions as possible.

Considering distributions with a **given mean and variance**, it can be proven (using variational inference) that such a distribution with **maximum entropy** is exactly the normal distribution.

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Maximum Likelihood Estimation

Let $X = \{x_1, x_2, \dots, x_N\}$ be training data drawn independently from the data-generating distribution p_{data} .

We denote the **empirical data distribution** as \hat{p}_{data} , where

$$\hat{p}_{ ext{data}}(oldsymbol{x}) \stackrel{ ext{\tiny def}}{=} rac{ig|\{i:oldsymbol{x}_i=oldsymbol{x}\}ig|}{N}$$

Let $p_{\text{model}}(\mathbf{x}; \boldsymbol{w})$ be a family of distributions.

- If the weights are fixed, $p_{ ext{model}}(\mathbf{x}; \boldsymbol{w})$ is a probability distribution.
- If we instead consider the fixed training data $oldsymbol{X}$, then

$$L(oldsymbol{w}) = p_{ ext{model}}(oldsymbol{X};oldsymbol{w}) = \prod_{i=1}^N p_{ ext{model}}(oldsymbol{x}_i;oldsymbol{w})$$

is called the **likelihood**. Note that even if the value of the likelihood is in range [0, 1], it is not a probability, because the likelihood is not a probability distribution.

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Maximum Likelihood Estimation

Let $X = \{x_1, x_2, \dots, x_N\}$ be training data drawn independently from the data-generating distribution p_{data} . We denote the empirical data distribution as \hat{p}_{data} and let $p_{\text{model}}(\mathbf{x}; \boldsymbol{w})$ be a family of distributions.

The maximum likelihood estimation of $oldsymbol{w}$ is:

$$egin{aligned} m{w}_{ ext{MLE}} &= rg\max_{m{w}} p_{ ext{model}}(m{X};m{w}) = rg\max_{m{w}} \prod_{i=1}^N p_{ ext{model}}(m{x}_i;m{w}) \ &= rg\min_{m{w}} \sum_{i=1}^N -\log p_{ ext{model}}(m{x}_i;m{w}) \ &= rg\min_{m{w}} \mathbb{E}_{m{x}\sim\hat{p}_{ ext{data}}}[-\log p_{ ext{model}}(m{x};m{w})] \ &= rg\min_{m{w}} H(\hat{p}_{ ext{data}}(m{x}),p_{ ext{model}}(m{x};m{w})) \ &= rg\min_{m{w}} D_{ ext{KL}}(\hat{p}_{ ext{data}}(m{x})\|p_{ ext{model}}(m{x};m{w})) + H(\hat{p}_{ ext{data}}(m{x})) \end{aligned}$$

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Maximum Likelihood Estimation

MLE can be easily generalized to the conditional case, where our goal is to predict t given \boldsymbol{x} :

$$egin{aligned} m{w}_{ ext{MLE}} &= rg\max_{m{w}} p_{ ext{model}}(m{t}|m{X};m{w}) = rg\max_{m{w}} \prod_{i=1}^N p_{ ext{model}}(t_i|m{x}_i;m{w}) \ &= rg\min_{m{w}} \sum_{i=1}^N -\log p_{ ext{model}}(t_i|m{x}_i;m{w}) \ &= rg\min_{m{w}} \mathbb{E}_{(\mathbf{x}, ext{t})\sim \hat{p}_{ ext{data}}}[-\log p_{ ext{model}}(m{t}|m{x};m{w})] \ &= rg\min_{m{w}} H(\hat{p}_{ ext{data}}(ext{t}|m{x}),p_{ ext{model}}(ext{t}|m{x};m{w})) \ &= rg\min_{m{w}} D_{ ext{KL}}(\hat{p}_{ ext{data}}(ext{t}|m{x})\|p_{ ext{model}}(ext{t}|m{x};m{w})) + H(\hat{p}_{ ext{data}}(ext{t}|m{x})) \end{aligned}$$

where the conditional entropy is defined as $H(\hat{p}_{\text{data}}) = \mathbb{E}_{(\mathbf{x},t) \sim \hat{p}_{\text{data}}}[-\log(\hat{p}_{\text{data}}(t|\mathbf{x}))]$ and the conditional cross-entropy as $H(\hat{p}_{\text{data}}, p_{\text{model}}) = \mathbb{E}_{(\mathbf{x},t) \sim \hat{p}_{\text{data}}}[-\log(p_{\text{model}}(t|\mathbf{x};\mathbf{w}))]$.

The resulting *loss function* is called **negative log-likelihood** (NLL), or **cross-entropy**, or **Kullback-Leibler divergence**.

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Logistic Regression



An extension of perceptron, which models the conditional probabilities of $p(C_0|\mathbf{x})$ and of $p(C_1|\mathbf{x})$. Logistic regression can in fact handle also more than two classes, which we will see in the next lecture.

Logistic regression employs the following parametrization of the conditional class probabilities:

$$egin{aligned} p(C_1 | oldsymbol{x}) &= \sigma(oldsymbol{x}^T oldsymbol{w} + b) \ p(C_0 | oldsymbol{x}) &= 1 - p(C_1 | oldsymbol{x}), \end{aligned}$$

where σ is a **sigmoid function**

$$\sigma(x)=rac{1}{1+e^{-x}}.$$

It can be trained using the SGD algorithm.

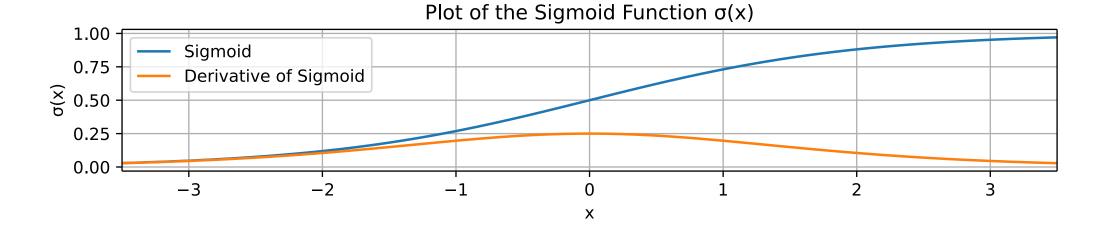
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Sigmoid Function



The sigmoid function has values in range (0, 1), is monotonically increasing and it has a derivative of $\frac{1}{4}$ at x = 0.

$$\sigma(x) = rac{1}{1+e^{-x}}
onumber \ \sigma'(x) = \sigma(x)ig(1-\sigma(x)ig)$$



Logistic Regression



We denote the output of the "linear part" of the logistic regression as

$$ar{y}(oldsymbol{x};oldsymbol{w})=oldsymbol{x}^Toldsymbol{w},$$

and the overall prediction as

$$y(oldsymbol{x};oldsymbol{w}) = \sigma(oldsymbol{ar{y}}(oldsymbol{x};oldsymbol{w})) = \sigma(oldsymbol{x}^Toldsymbol{w}).$$

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Logistic Regression

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To train the logistic regression, we use MLE (the maximum likelihood estimation). Its application is straightforward, given that $p(C_1|\boldsymbol{x}; \boldsymbol{w})$ is directly the model output $y(\boldsymbol{x}; \boldsymbol{w})$.

Therefore, the loss for a minibatch $\mathbb{X} = \{(\boldsymbol{x}_1, t_1), (\boldsymbol{x}_2, t_2), \dots, (\boldsymbol{x}_N, t_N)\}$ is

$$E(oldsymbol{w}) = rac{1}{N}\sum_i -\log(p(C_{t_i}|oldsymbol{x}_i;oldsymbol{w})).$$

Input: Input dataset ($m{X} \in \mathbb{R}^{N imes D}$, $m{t} \in \{0,+1\}^N$), learning rate $lpha \in \mathbb{R}^+$.

- $oldsymbol{w} \leftarrow oldsymbol{0}$ or we initialize $oldsymbol{w}$ randomly
- until convergence (or patience runs out), process a minibatch of examples \mathbb{B} :

$$\circ ~~ oldsymbol{g} \leftarrow rac{1}{|\mathbb{B}|} \sum_{i \in \mathbb{B}}
abla_{oldsymbol{w}} \Big(-\logig(p(C_{t_i} | oldsymbol{x}_i; oldsymbol{w}) ig) \Big)$$

 $\circ \boldsymbol{w} \leftarrow \boldsymbol{w} - lpha \boldsymbol{g}$



Everything we learned about **features** and L^2 **regularization** holds for logistic regression too.



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