

# Introduction to Introduction to Natural Language Processing / Úvod do zpracování přirozeného jazyka

Lekce 1-2

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# Intro to NLP

- Instructor: Jan Hajič, Pavel Pecina
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  - Hours: J. Hajic: Mon 10:00-11:00
  - preferred contact:  [{hajic,pecina}@ufal.mff.cuni.cz](mailto:{hajic,pecina}@ufal.mff.cuni.cz)
- Room & time:
  - lecture: SU1, Wed, 15:40-17:10 + S7, Wed, 17:20-18:50
  - seminar [cvičení] follows (Pavel Pecina, Zdeněk Žabokrtský, ...)
  - Other info: pls see at the seminar

# Textbooks you need

- Manning, C. D., Schütze, H.:
  - *Foundations of Statistical Natural Language Processing*. The MIT Press. 1999. ISBN 0-262-13360-1. **[available at least at MFF / Computer Science School library, Malostranske nam. 25, 11800 Prague 1]**
- Jurafsky, D., Martin, J.H.:
  - *Speech and Language Processing*. Prentice-Hall. 2000. ISBN 0-13-095069-6 and **newer editions**. **[recommended]**.
- Cover, T. M., Thomas, J. A.:
  - *Elements of Information Theory*. Wiley. 1991. ISBN 0-471-06259-6.
- Jelinek, F.:
  - *Statistical Methods for Speech Recognition*. The MIT Press. 1998. ISBN 0-262-10066-5

# Other reading

- Journals:
  - Computational Linguistics
  - Transactions on Computational Linguistics
- Proceedings of major conferences:
  - ACL (Assoc. of Computational Linguistics)
  - EACL (European Chapter of ACL)
  - EMNLP (Empirical Methods in NLP)
  - CoNLL (Natural Language Learning in CL)
  - IJCNLP (Asian chapter of ACL)
  - COLING (Intl. Committee of Computational Linguistics)

# Course segments (first three lectures)

- Intro & Probability & Information Theory
  - The very basics: definitions, formulas, examples.
- Language Modeling
  - n-gram models, parameter estimation
  - smoothing (EM algorithm)

# Probability

# Experiments & Sample Spaces

- Experiment, process, test, ...
- Set of possible basic outcomes: sample space  $\Omega$ 
  - coin toss ( $\Omega = \{\text{head}, \text{tail}\}$ ), die ( $\Omega = \{1..6\}$ )
  - yes/no opinion poll, quality test (bad/good) ( $\Omega = \{0,1\}$ )
  - lottery ( $|\Omega| \cong 10^7 \dots 10^{12}$ )
  - # of traffic accidents somewhere per year ( $\Omega = \mathbb{N}$ )
  - spelling errors ( $\Omega = Z^*$ ), where  $Z$  is an alphabet, and  $Z^*$  is a set of possible strings over such an alphabet
  - missing word ( $|\Omega| \cong \text{vocabulary size}$ )

# Events

- Event  $A$  is a set of basic outcomes
- Usually  $A \subset \Omega$ , and all  $A \in 2^\Omega$  (the event space)
  - $\Omega$  is then the certain event,  $\emptyset$  is the impossible event
- Example:
  - experiment: three times coin toss
    - $\Omega = \{\mathbf{HHH}, \mathbf{HHT}, \mathbf{HTH}, \mathbf{HTT}, \mathbf{THH}, \mathbf{THT}, \mathbf{TTH}, \mathbf{TTT}\}$
  - count cases with exactly two tails: then
    - $A = \{\mathbf{HTT}, \mathbf{THT}, \mathbf{TTH}\}$
  - all heads:
    - $A = \{\mathbf{HHH}\}$



# Probability

- Repeat experiment many times, record how many times a given event  $A$  occurred (“count”  $c_1$ ).
- Do this whole series many times; remember all  $c_i$ s.
- Observation: if repeated really many times, the ratios of  $c_i/T_i$  (where  $T_i$  is the number of experiments run in the  $i$ -th series) are close to some (unknown but) **constant** value.
- Call this constant a **probability of  $A$** . Notation:  **$p(A)$**

# Estimating probability

- Remember: ... close to an *unknown* constant.
- We can only estimate it:
  - from a single series (typical case, as mostly the outcome of a series is given to us and we cannot repeat the experiment), set
$$p(A) = c_1/T_1.$$
  - otherwise, take the weighted average of all  $c_i/T_i$  (or, if the data allows, simply look at the set of series as if it is a single long series).
- This is the **best** estimate.

# Example

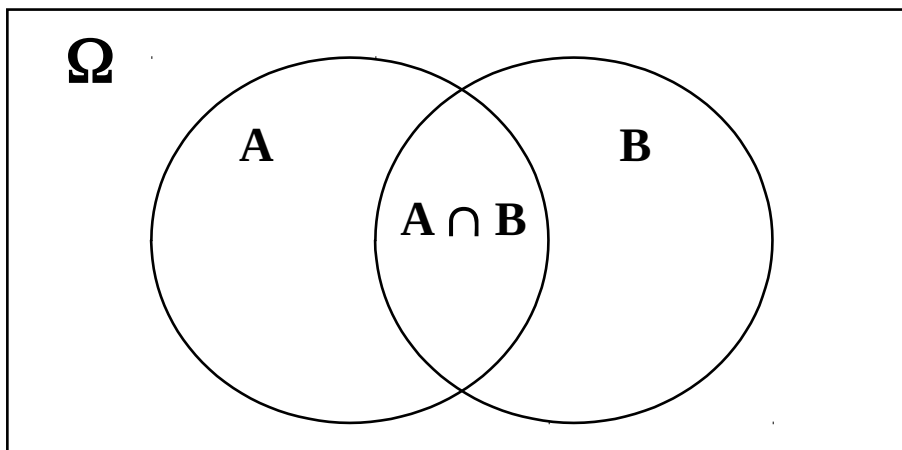
- Recall our example:
  - experiment: three times coin toss
    - $\Omega = \{\text{HHH}, \text{HHT}, \text{HTH}, \text{HTT}, \text{TTH}, \text{THT}, \text{TTH}, \text{TTT}\}$
  - count cases with exactly two tails:  $A = \{\text{HTT}, \text{THT}, \text{TTH}\}$
- Run an experiment 1000 times (i.e. 3000 tosses)
- Counted: 386 cases with two tails ( $\text{HTT}$ ,  $\text{THT}$ , or  $\text{TTH}$ )
- estimate:  $p(A) = 386 / 1000 = .386$
- Run again: 373, 399, 382, 355, 372, 406, 359
  - $p(A) = .379$  (weighted average) or simply  $3032 / 8000$
- *Uniform* distribution assumption:  $p(A) = 3/8 = .375$

# Basic Properties

- Basic properties:
  - $p: 2^\Omega \rightarrow [0,1]$
  - $p(\Omega) = 1$
  - Disjoint events:  $p(\cup A_i) = \sum_i p(A_i)$
- [NB: axiomatic definition of probability: take the above three conditions as axioms]
- Immediate consequences:
  - $p(\emptyset) = 0$ ,  $p(\bar{A}) = 1 - p(A)$ ,  $A \subseteq B \Rightarrow p(A) \leq p(B)$
  - $\sum_{a \in \Omega} p(a) = 1$

# Joint and Conditional Probability

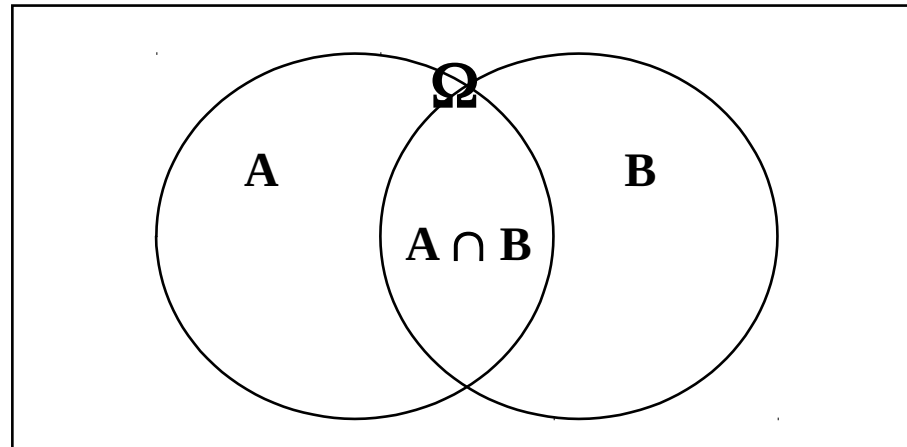
- $p(A,B) = p(A \cap B)$
- $p(A|B) = p(A,B) / p(B)$ 
  - Estimating from counts:
    - $p(A|B) = p(A,B) / p(B) = (c(A \cap B) / T) / (c(B) / T) = c(A \cap B) / c(B)$



# Bayes Rule

- $p(A,B) = p(B,A)$  since  $p(A \cap B) = p(B \cap A)$ 
  - therefore:  $p(A|B) p(B) = p(B|A) p(A)$ , and therefore

$$p(A|B) = p(B|A) p(A) / p(B)$$



# Independence

- Can we compute  $p(A,B)$  from  $p(A)$  and  $p(B)$ ?
- Recall from previous foil:

$$p(A|B) = p(B|A) p(A) / p(B)$$

$$p(A|B) p(B) = p(B|A) p(A)$$

$$p(A,B) = p(B|A) p(A)$$

... we're almost there: how  $p(B|A)$  relates to  $p(B)$ ?

–  $p(B|A) = P(B)$  (iff) A and B are **independent**

- Example: two coin tosses, weather today and weather on March 4th 1789;
- Any two events for which  $p(B|A) = P(B)$ !

# Chain Rule

$$p(A_1, A_2, A_3, A_4, \dots, A_n) =$$



$$p(A_1|A_2, A_3, A_4, \dots, A_n) \times p(A_2|A_3, A_4, \dots, A_n) \times \\ \times p(A_3|A_4, \dots, A_n) \times \dots p(A_{n-1}|A_n) \times p(A_n)$$

- this is a direct consequence of the Bayes rule.



# The Golden Rule (of Classic Statistical NLP)

- Interested in an event A given B (when it is not easy or practical or desirable to estimate  $p(A|B)$ ):
- take Bayes rule, max over all As:
- $\operatorname{argmax}_A p(A|B) = \operatorname{argmax}_A p(B|A) \cdot p(A) / p(B) =$

$\operatorname{argmax}_A p(B|A) p(A) \quad \bullet$

- ... as  $p(B)$  is constant when changing As

# Random Variable

- is a function  $X: \Omega \rightarrow Q$ 
  - in general:  $Q = \mathbb{R}^n$ , typically  $\mathbb{R}$
  - easier to handle real numbers than real-world events
- random variable is *discrete* if  $Q$  is countable (i.e. also if finite)
- Example: *die*: natural “numbering”  $[1,6]$ , *coin*:  $\{0,1\}$
- Probability distribution:
  - $p_X(x) = p(X=x) =_{\text{df}} p(A_x)$  where  $A_x = \{a \in \Omega : X(a) = x\}$
  - often just  $p(x)$  if it is clear from context what  $X$  is

# Expectation

## Joint and Conditional Distributions

- is a mean of a random variable (weighted average)
  - $E(X) = \sum_{x \in X(\Omega)} x \cdot p_X(x)$
- Example: one six-sided die: 3.5, two dice (sum) 7
- Joint and Conditional distribution rules:
  - analogous to probability of events
- Bayes:  $p_{X|Y}(x,y) =$  notation  $p_{XY}(x|y) =$  even simpler notation  
 **$p(x|y) = p(y|x) \cdot p(x) / p(y)$**
- Chain rule:  **$p(w,x,y,z) = p(z) \cdot p(y|z) \cdot p(x|y,z) \cdot p(w|x,y,z)$**

# Essential Information Theory

# The Notion of Entropy

- Entropy ~ “chaos”, fuzziness, opposite of order, ...
  - you know it:
    - **it is much easier to create “mess” than to tidy things up...**
- Comes from physics:
  - Entropy does not go down unless energy is applied
- Measure of **uncertainty**:
  - if low... low uncertainty; the higher the entropy, the higher uncertainty, but the higher “surprise” (information) we can get out of an experiment

# The Formula

- Let  $p_X(x)$  be a distribution of random variable  $X$
- Basic outcomes (alphabet)  $\Omega$

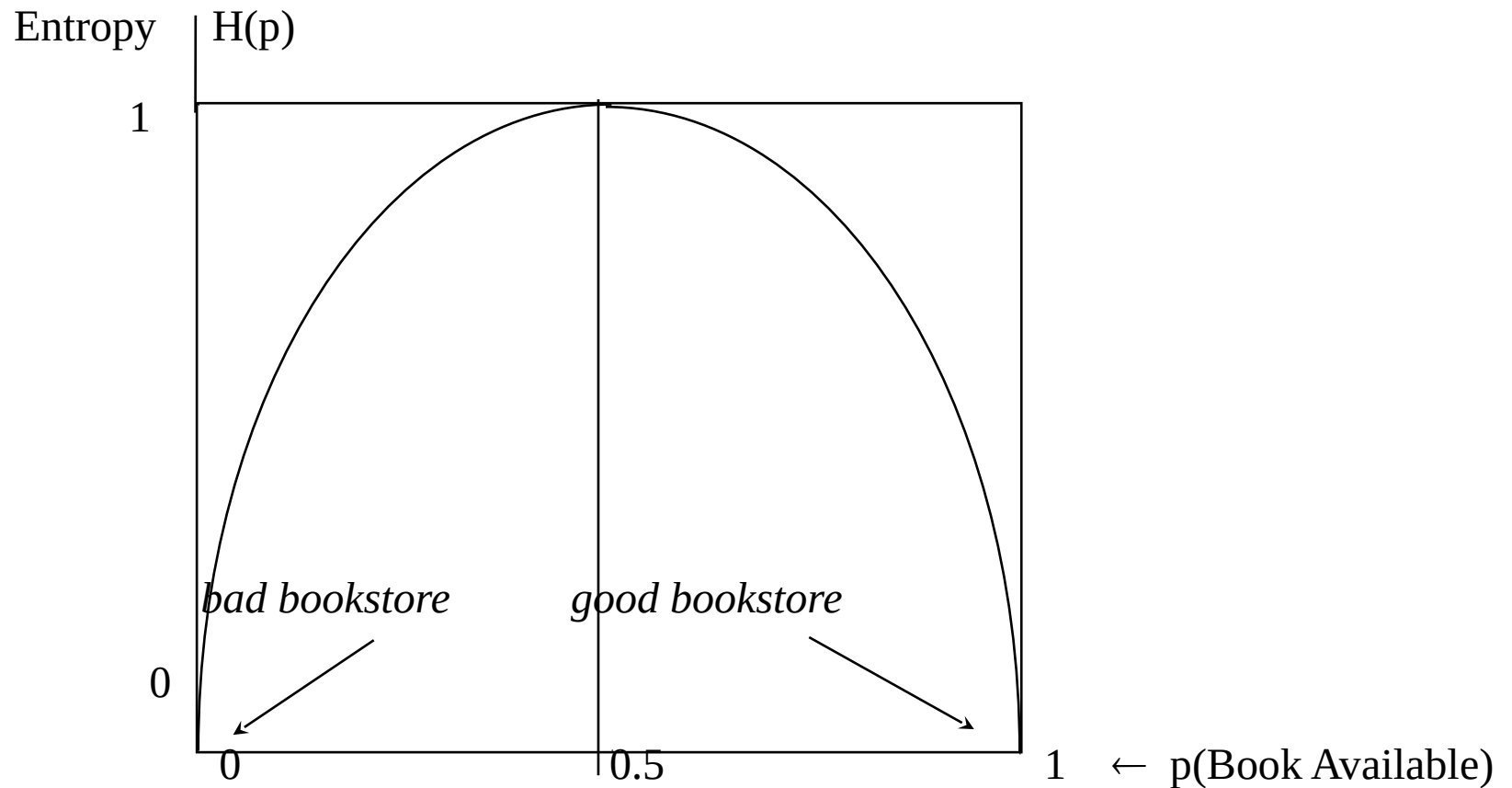
$$H(X) = - \sum_{x \in \Omega} p(x) \log_2 p(x) \quad !$$

- Unit: bits ( $\log_{10}$ : nats)
- Notation:  $H(X) = H_p(X) = H(p) = H_X(p) = H(p_X)$

# Using the Formula: Example

- Toss a fair coin:  $\Omega = \{\text{head}, \text{tail}\}$ 
  - $p(\text{head}) = .5, p(\text{tail}) = .5$
  - $\mathbf{H(p)} = -0.5 \log_2(0.5) + (-0.5 \log_2(0.5)) = 2 \times ((-0.5) \times (-1)) = 2 \times 0.5 = \mathbf{1}$
- Take fair, 32-sided die:  $p(x) = 1 / 32$  for every side  $x$ 
  - $\mathbf{H(p)} = -\sum_{i=1..32} p(x_i) \log_2 p(x_i) = -32 (p(x_1) \log_2 p(x_1))$   
(since for all  $i$   $p(x_i) = p(x_1) = 1/32$ )  
 $= -32 \times ((1/32) \times (-5)) = \mathbf{5}$  (now you see why it's called **bits**?)
- Unfair coin:
  - $p(\text{head}) = .2 \dots \mathbf{H(p)} = .722$ ;  $p(\text{head}) = .01 \dots \mathbf{H(p)} = .081$

# Example: Book Availability





# The Limits

- When  $H(p) = 0$ ?
  - if a result of an experiment is *known* ahead of time:
  - necessarily:
$$\exists x \in \Omega; p(x) = 1 \ \& \ \forall y \in \Omega; y \neq x \Rightarrow p(y) = 0$$
- Upper bound?
  - none in general
  - for  $|\Omega| = n$ :  $H(p) \leq \log_2 n$ 
    - **nothing can be more uncertain than the uniform distribution**

# Perplexity: motivation

- Recall:
  - 2 equiprobable outcomes:  $H(p) = 1$  bit
  - 32 equiprobable outcomes:  $H(p) = 5$  bits
  - 4.3 billion equiprobable outcomes:  $H(p) \approx 32$  bits
- What if the outcomes are not equiprobable?
  - 32 outcomes, 2 equiprobable at .5, rest impossible:
    - **$H(p) = 1$  bit**
  - Any measure for comparing the entropy (i.e. uncertainty/difficulty of prediction) (also) for random variables with different number of outcomes?

# Perplexity

- Perplexity:
  - $G(p) = 2^{H(p)}$
- ... so we are back at 32 (for 32 eqp. outcomes), 2 for fair coins, etc.
- it is easier to imagine:
  - NLP example: vocabulary size of a vocabulary with uniform distribution, which is equally hard to predict
- the “wilder” (biased) distribution, the better:
  - lower entropy, lower perplexity

# Joint Entropy and Conditional Entropy

- Two random variables:  $X$  (space  $\Omega$ ),  $Y$  ( $\Psi$ )
- Joint entropy:
  - no big deal:  $((X, Y)$  considered a single event):

$$H(X, Y) = - \sum_{x \in \Omega} \sum_{y \in \Psi} p(x, y) \log_2 p(x, y)$$

- Conditional entropy:

$$H(Y|X) = - \sum_{x \in \Omega} \sum_{y \in \Psi} \underline{p(x, y)} \log_2 p(y|x)$$

recall that  $H(X) = E(\log_2(1/p_X(x)))$

(weighted average: weights are not conditional)

# Properties of Entropy I

- Entropy is non-negative:
  - $H(X) \geq 0$
  - proof: (recall:  $H(X) = - \sum_{x \in \Omega} p(x) \log_2 p(x)$ )
    - **$\log(p(x))$  is negative or zero for  $x \leq 1$ ,**
    - **$p(x)$  is non-negative; their product  $p(x)\log(p(x))$  is thus negative;**
    - **sum of negative numbers is negative;**
    - **and  $-f$  is positive for negative  $f$**
- Chain rule:
  - $H(X, Y) = H(Y|X) + H(X)$ , as well as
  - $H(X, Y) = H(X|Y) + H(Y)$  (since  $H(Y, X) = H(X, Y)$ )

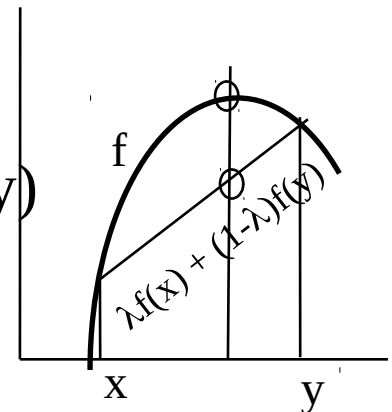
# Properties of Entropy II

- Conditional Entropy is better (than unconditional):
  - $H(Y|X) \leq H(Y)$
- $H(X, Y) \leq H(X) + H(Y)$  (follows from the previous (in)equalities)
  - **equality iff  $X, Y$  independent**
  - **[recall:  $X, Y$  independent iff  $p(X, Y) = p(X)p(Y)$ ]**
- $H(p)$  is concave (remember the book availability graph?)
  - concave function  $f$  over an interval  $(a, b)$ :

$$\forall x, y \in (a, b), \forall \lambda \in [0, 1]:$$

$$f(\lambda x + (1-\lambda)y) \geq \lambda f(x) + (1-\lambda)f(y)$$

- **function  $f$  is convex if  $-f$  is concave**



# “Coding” Interpretation of Entropy

- The least (average) number of bits needed to encode a message (string, sequence, series,...) (each element having being a result of a random process with some distribution  $p$ ):  $= H(p)$
- Remember various compressing algorithms?
  - they do well on data with repeating (= easily predictable = low entropy) patterns
  - their results though have high entropy  $\Rightarrow$  compressing compressed data does nothing

# Coding: Example

- How many bits do we need for ISO Latin 1?
  - $\Rightarrow$  the trivial answer: 8
- Experience: some chars are more common, some (very) rare:
  - ...so what if we use more bits for the rare, and less bits for the frequent? [be careful: want to decode (easily)!] **!**
  - suppose:  $p('a') = 0.3$ ,  $p('b') = 0.3$ ,  $p('c') = 0.3$ , the rest:  $p(x) \cong .0004$
  - code: 'a'  $\sim$  00, 'b'  $\sim$  01, 'c'  $\sim$  10, rest:  $11b_1b_2b_3b_4b_5b_6b_7b_8$
  - code acbbécbaac: **001001011110000111111001000010**  
  **a c b b                    é                    c b a a c**
  - number of bits used: 28 (vs. 80 using “naive” coding)
- code length  $\sim 1 / \text{probability}$ ; conditional prob OK!



# Kullback-Leibler Distance (Relative Entropy)

- Remember:
  - long series of experiments...  $c_i/T_i$  oscillates around some number... we can only estimate it... to get a distribution  $\underline{q}$ .
- So we get a distribution  $\underline{q}$ ; (sample space  $\Omega$ , r.v.  $X$ )  
the true distribution is, however,  $\underline{p}$ . (same  $\Omega$ ,  $X$ )  
 $\Rightarrow$  how big error are we making?
- $D(p||q)$  (the Kullback-Leibler distance):

$$D(p||q) = \sum_{x \in \Omega} \underline{p}(x) \log_2 (p(x)/q(x)) = E_p \log_2 (p(x)/q(x))$$

# Comments on Relative Entropy

- Conventions:
  - $0 \log 0 = 0$
  - $p \log (p/0) = \infty$  (for  $p > 0$ )
- Distance? (less “misleading”: Divergence)
  - not quite:
    - **not symmetric:  $D(p||q) \neq D(q||p)$**
    - **does not satisfy the triangle inequality**
  - but useful to look at it that way
- $H(p) + D(p||q)$ : bits needed for encoding  $p$  if  $q$  is used

# Mutual Information (MI) in terms of relative entropy

- Random variables  $X, Y$ ;  $p_{X \cap Y}(x,y)$ ,  $p_X(x)$ ,  $p_Y(y)$
- Mutual information (between two random variables  $X, Y$ ):


$$I(X, Y) = D(p(x,y) \parallel p(x)p(y))$$

- $I(X, Y)$  measures how much (our knowledge of)  $Y$  contributes (on average) to easing the prediction of  $X$
- or, how  $p(x,y)$  deviates from (independent)  $p(x)p(y)$

# Mutual Information: the Formula

- Rewrite the definition: [recall:  $D(r||s) = \sum_{v \in \Omega} r(v) \log_2 (r(v)/s(v))$ ];  
substitute  $r(v) = p(x,y)$ ,  $s(v) = p(x)p(y)$ ;  $\langle v \rangle \sim$

$\langle x,y \rangle$

$$I(X, Y) = D(p(x,y) || p(x)p(y)) =$$
$$= \sum_{x \in \Omega} \sum_{y \in \Psi} p(x,y) \log_2 (p(x,y)/p(x)p(y))$$


- Measured in bits (what else? :-)

# From Mutual Information to Entropy

- by how many bits the knowledge of  $Y$  **lowers** the entropy  $H(X)$ :

$$\begin{aligned}
 I(X, Y) &= \sum_{x \in \Omega} \sum_{y \in \Psi} p(x, y) \log_2 \left( \frac{p(x, y)}{p(y)p(x)} \right) = \\
 &\quad \dots \text{use } p(x, y)/p(y) = p(x|y) \\
 &= \sum_{x \in \Omega} \sum_{y \in \Psi} p(x, y) \log_2 \left( \frac{p(x|y)}{p(x)} \right) = \\
 &\quad \dots \text{use } \log(a/b) = \log a - \log b \text{ (} a \sim p(x|y), b \sim p(x)\text{), distribute sums} \\
 &= \frac{\sum_{x \in \Omega} \sum_{y \in \Psi} p(x, y) \log_2 p(x|y)}{\dots \text{use def. of } H(X|Y) \text{ (left term), and } \sum_{y \in \Psi} p(x, y) = p(x) \text{ (right term)}} - \sum_{x \in \Omega} \sum_{y \in \Psi} p(x, y) \log_2 p(x) = \\
 &= \frac{-H(X|Y)}{\dots \text{use def. of } H(X) \text{ (right term), swap terms}} + \left( - \sum_{x \in \Omega} p(x) \log_2 p(x) \right) = \\
 &= H(X) - H(X|Y) \quad \dots \text{by symmetry, } = H(Y) - H(Y|X)
 \end{aligned}$$

# Properties of MI vs. Entropy

- $I(X, Y) = H(X) - \underbrace{H(X|Y)}_{\text{of Y lowers the entropy of X}} = \text{number of bits the knowledge of Y lowers the entropy of X}$   
 $= H(Y) - H(Y|X)$  (prev. foil, symmetry)

Recall:  $H(X, Y) = H(X|Y) + H(Y) \Rightarrow H(X|Y) = H(Y) - H(X, Y) \Rightarrow$

- $I(X, Y) = H(X) + \underline{H(Y) - H(X, Y)}$
- $I(X, X) = H(X)$  (since  $H(X|X) = 0$ )
- $I(X, Y) = I(Y, X)$  (just for completeness)
- $I(X, Y) \geq 0$  ... let's prove that now (as promised).

# Other (In)Equalities and Facts

- Log sum inequality: for  $r_i, s_i \geq 0$

$$\sum_{i=1..n} (r_i \log(r_i/s_i)) \leq \left(\sum_{i=1..n} r_i\right) \log\left(\sum_{i=1..n} r_i / \sum_{i=1..n} s_i\right)$$

- $D(p||q)$  is convex [in  $p, q$ ] ( $\Leftarrow$  log sum inequality)
- $H(p_X) \leq \log_2|\Omega|$ , where  $\Omega$  is the sample space of  $p_X$

Proof: uniform  $u(x)$ , same sample space  $\Omega$ :  $\sum p(x) \log u(x) = -\log_2|\Omega|$ ;

$$\log_2|\Omega| - H(X) = -\sum p(x) \log u(x) + \sum p(x) \log p(x) = D(p||u) \geq 0$$

- $H(p)$  is concave [in  $p$ ]:

Proof: from  $H(X) = \log_2|\Omega| - D(p||u)$ ,  $D(p||u)$  convex  $\Rightarrow H(x)$  concave

# Cross-Entropy

- Typical case: we've got series of observations  
 $T = \{t_1, t_2, t_3, t_4, \dots, t_n\}$  (numbers, words, ...;  $t_i \in \Omega$ );  
estimate (simple):  
 $\forall y \in \Omega: \tilde{p}(y) = c(y) / |T|$ , def.  $c(y) = |\{t \in T; t = y\}|$
- ...but the true  $p$  is unknown; every sample is too small!
- Natural question: how well do we do using  $\tilde{p}$  [instead of  $p$ ]?
- Idea: simulate actual  $p$  by using a different  $T'$   
(or rather: by using different observation we simulate the insufficiency of  $T$  vs. some other data (“random” difference))



# Cross Entropy: The Formula

- $H_{p'}(\tilde{p}) = H(p') + D(p' \| \tilde{p})$

$$H_{p'}(\tilde{p}) = - \sum_{x \in \mathcal{O}} p'(x) \log_{\tilde{p}}(x) \quad \text{!}$$

- $p'$  is certainly not the true  $p$ , but we can consider it the “real world” distribution against which we test  $\tilde{p}$
- note on notation (confusing...):  $p/p' \leftrightarrow \tilde{p}$ , also  $H_{T'}(p)$
- (Cross)Perplexity:  $G_{p'}(p) = G_{T'}(p) = 2^{H_{p'}(p)}$

# Conditional Cross Entropy

- So far: “unconditional” distribution(s)  $p(x)$ ,  $p'(x)$ ...
- In practice: virtually always conditioning on context
- Interested in: sample space  $\Psi$ , r.v.  $Y$ ,  $y \in \Psi$ ;  
context: sample space  $\Omega$ , r.v.  $X$ ,  $x \in \Omega$ ;;  
“our” distribution  $p(y|x)$ , test against  
 $p'(y,x)$ ,  
which is taken from some independent data:

$$H_{p'}(p) = - \sum_{y \in \Psi, x \in \Omega} p'(y,x) \log_2 p(y|x)$$

# Sample Space vs. Data

- In practice, it is often inconvenient to sum over the sample space(s)  $\Psi, \Omega$  (especially for cross entropy!)
- Use the following formula:

$$H_{p'}(p) = - \sum_{y \in \Psi, x \in \Omega} p'(y,x) \log_2 p(y|x) = - 1/|T'| \sum_{i=1..|T'|} \log_2 p(y_i|x_i) \quad !$$

- This is in fact the normalized log probability of the “test” data:

$$H_{p'}(p) = - 1/|T'| \log_2 \prod_{i=1..|T'|} p(y_i|x_i)$$

# Computation Example

- $\Omega = \{a,b,\dots,z\}$ , prob. distribution (assumed/estimated from data):

$p(a) = .25, p(b) = .5, p(\alpha) = 1/64$  for  $\alpha \in \{c..r\}, = 0$  for the rest:  
 $s,t,u,v,w,x,y,z$

- Data (test): barb  $p'(a) = p'(r) = .25, p'(b) = .5$

- Sum over  $\Omega$ :

$\alpha$	<b>a</b>	<b>b</b>	<b>c</b>	<b>d</b>	<b>e</b>	<b>f</b>	<b>g</b>	<b>...</b>	<b>p</b>	<b>q</b>	<b>r</b>	<b>s</b>	<b>t</b>	<b>...</b>	<b>z</b>
$-p'(\alpha)\log_2 p(\alpha)$	<b>.5</b>	<b>+ .5</b>	<b>+ 0</b>	<b>+ 0</b>	<b>+ 0</b>	<b>+ 0</b>	<b>+ 0</b>	<b>+ 0</b>	<b>+ 0</b>	<b>+ 0</b>	<b>+ 0</b>	<b>+ 0</b>	<b>+ 0</b>	<b>+ 0</b>	<b>+ 1.5</b>
	<b>= <u>2.5</u></b>														

- Sum over data:

$i / s_i$	<b>1/b</b>	<b>2/a</b>	<b>3/r</b>	<b>4/b</b>	<b>= 10</b>	$\left(\frac{1}{4}\right) \times$	<b>10</b>	<b>=</b>
$-\log_2 p(s_i)$	<b>1</b>	<b>2</b>	<b>6</b>	<b>1</b>	<b>= 10</b>	$\left(\frac{1}{4}\right) \times$	<b>10</b>	<b>=</b>
	<b><u>2.5</u></b>							

↖  $1/|T'|$

# Cross Entropy: Some Observations

- $H(p) \quad ?? <, =, > ?? \quad H_{p'}(p): \text{ALL!}$

- Previous example:

$[p(a) = .25, p(b) = .5, p(\alpha) = 1/64 \text{ for } \alpha \in \{c..r\}, = 0 \text{ for the rest: } s,t,u,v,w,x,y,z]$

$$H(p) = 2.5 \text{ bits} = H(p') \text{ (barb)}$$

- Other data: probable:  $(1/8)(6+6+6+1+2+1+6+6) = 4.25$

$$H(p) < 4.25 \text{ bits} = H(p') \text{ (probable)}$$

- And finally: abba:  $(1/4)(2+1+1+2) = 1.5$

$$H(p) > 1.5 \text{ bits} = H(p') \text{ (abba)}$$

- But what about: baby  $-p'('y')\log_2 p('y') = -.25\log_2 0 = \infty$   
(??)

# Cross Entropy: Usage

- Comparing data??
  - NO! (we believe that we test on real data!)
- Rather: comparing distributions (vs. real data)
- Have (got) 2 distributions:  $p$  and  $q$  (on some  $\Omega$ ,  $X$ )
  - which is better?
  - better: has lower cross-entropy (perplexity) on real data  $S$
- “Real” data:  $S$
- $H_S(p) = -1/|S| \sum_{i=1..|S|} \log_2 p(y_i|x_i)$ . ??  $H_S(q) = -1/|S| \sum_{i=1..|S|} \log_2 q(y_i|x_i)$

# Comparing Distributions

Test data S: probable

- $p(\cdot)$  from prev. example:

$$H_S(p) = 4.25$$

$p(a) = .25, p(b) = .5, p(\alpha) = 1/64$  for  $\alpha \in \{c..r\}, = 0$  for the rest: s,t,u,v,w,x,y,z

- $q(\cdot|\cdot)$  (conditional; defined by a table):

$q(\cdot \cdot) \begin{matrix} \rightarrow \\ \downarrow \end{matrix}$	a	b	e	l	o	p	r	other
a	0	.5	0	0	0	.125	0	0
b	1	0	0	0	1	.125	0	0
e	0	0	0	1	0	.125	0	0
l	0	.5	0	0	0	.125	0	0
o	0	0	0	0	0	.125	1	0
p	0	0	0	0	0	.125	0	1
r	0	0	0	0	0	.125	0	0
other	0	0	1	0	0	.125	0	0

ex.:  $q(o|r) = 1$

$q(r|p) = .125$

$$(1/8) (\log(p|oth.) + \log(r|p) + \log(o|r) + \log(b|o) + \log(a|b) + \log(b|a) + \log(l|b) + \log(e|l))$$

$$(1/8) ( 0 + 3 + 0 + 0 + 1 + 0 + 1 + 0 )$$

$$H_S(q) = .625$$

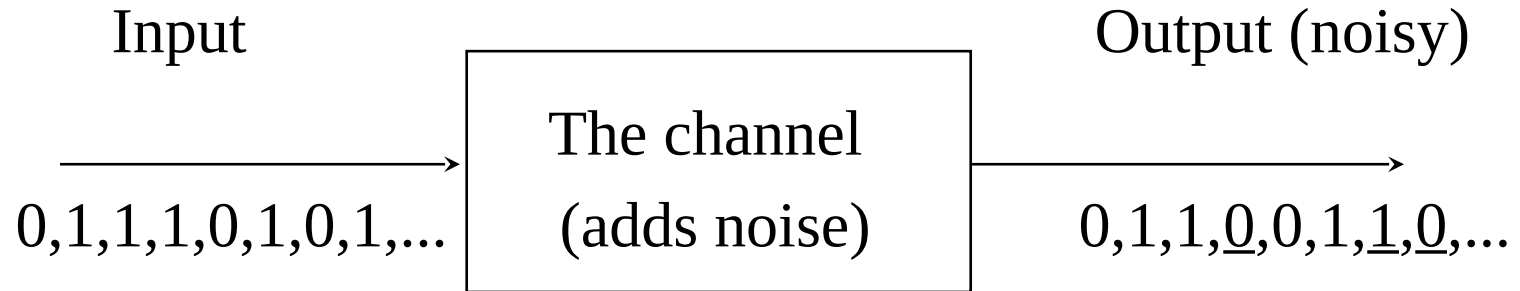


# Language Modeling (and the Noisy Channel)



# The Noisy Channel

- Prototypical case:



- Model: probability of error (noise):
- Example:  $p(0|1) = .3$   $p(1|1) = .7$   $p(1|0) = .4$   $p(0|0) = .6$
- The Task:  
known: the noisy output; want to know: the input (***decoding***)

# Noisy Channel Applications

- OCR
  - straightforward: text → print (adds noise), scan → image
- Handwriting recognition
  - text → neurons, muscles (“noise”), scan/digitize → image
- Speech recognition (dictation, commands, etc.)
  - text → conversion to acoustic signal (“noise”) → acoustic waves
- Machine Translation
  - text in target language → translation (“noise”) → source language
- Also: Part of Speech Tagging
  - sequence of tags → selection of word forms → text

# Noisy Channel: The Golden Rule of ...

OCR, ASR, HR, MT, ...

- Recall:

$$p(A|B) = p(B|A) p(A) / p(B) \quad (\text{Bayes formula})$$

$$A_{\text{best}} = \operatorname{argmax}_A p(B|A) p(A) \quad (\text{The Golden Rule})$$

- $p(B|A)$ : the acoustic/image/translation/lexical model
  - application-specific name
  - will explore later
- $p(A)$ : **the language model**

# The Perfect Language Model

- Sequence of word forms [forget about tagging for the moment]
- Notation:  $A \sim W = (w_1, w_2, w_3, \dots, w_d)$

- The big (modeling) question:

$$p(W) = ?$$

- Well, we know (Bayes/chain rule  $\rightarrow$ ):

$$\begin{aligned} p(W) &= p(w_1, w_2, w_3, \dots, w_d) = \\ &= p(w_1) \times p(w_2|w_1) \times p(w_3|w_1, w_2) \times \dots \times p(w_d| \\ & \quad w_1, w_2, \dots, w_{d-1}) \end{aligned}$$

- Not practical (even short  $W \rightarrow$  too many parameters)

# Markov Chain

- Unlimited memory (cf. previous foil):
  - for  $w_i$ , we know all its predecessors  $w_1, w_2, w_3, \dots, w_{i-1}$
- Limited memory:
  - we disregard “too old” predecessors
  - remember only  $k$  previous words:  $w_{i-k}, w_{i-k+1}, \dots, w_{i-1}$
  - called “ $k^{\text{th}}$  order Markov approximation”
- + stationary character (no change over time):

$$p(W) \cong \prod_{i=1..d} p(w_i | w_{i-k}, w_{i-k+1}, \dots, w_{i-1}), \quad d = |W|$$

# n-gram Language Models

- $(n-1)^{\text{th}}$  order Markov approximation  $\rightarrow$  n-gram LM:

prediction history

$p(W) =_{\text{df}} \prod_{i=1..d} p(w_i | w_{i-n+1}, w_{i-n+2}, \dots, w_{i-1})$  !

- In particular (assume vocabulary  $|V| = 60\text{k}$ ):
  - **0-gram LM: uniform model**,  $p(w) = 1/|V|$ , 1 parameter
  - **1-gram LM: unigram model**,  $p(w)$ ,  $6 \times 10^4$  parameters
  - **2-gram LM: bigram model**,  $p(w_i | w_{i-1})$   $3.6 \times 10^9$  parameters
  - **3-gram LM: trigram model**,  $p(w_i | w_{i-2}, w_{i-1})$   $2.16 \times 10^{14}$  parameters

# Maximum Likelihood Estimate

- MLE: Relative Frequency...
  - ...best predicts the data at hand (the “training data”)
- Trigrams from Training Data T:
  - count sequences of three words in T:  $c_3(w_{i-2}, w_{i-1}, w_i)$   
[NB: notation: just saying that the three words follow each other]
  - count sequences of two words in T:  $c_2(w_{i-1}, w_i)$ :
    - **either use  $c_2(y,z) = \sum_w c_3(y,z,w)$**
    - **or count differently at the beginning (& end) of data!**

$$p(w_i | w_{i-2}, w_{i-1}) =_{\text{est.}} c_3(w_{i-2}, w_{i-1}, w_i) / c_2(w_{i-2}, w_{i-1}) \quad !$$

# LM: an Example

- Training data:

$\langle s \rangle \langle s \rangle$  He can buy the can of soda.

- Unigram:  $p_1(\text{He}) = p_1(\text{buy}) = p_1(\text{the}) = p_1(\text{of}) = p_1(\text{soda}) = p_1(\cdot) = .125$

$p_1(\text{can}) = .25$

- Bigram:  $p_2(\text{He}|\langle s \rangle) = 1$ ,  $p_2(\text{can}|\text{He}) = 1$ ,  $p_2(\text{buy}|\text{can}) = .5$ ,

$p_2(\text{of}|\text{can}) = .5$ ,  $p_2(\text{the}|\text{buy}) = 1, \dots$

- Trigram:  $p_3(\text{He}|\langle s \rangle, \langle s \rangle) = 1$ ,  $p_3(\text{can}|\langle s \rangle, \text{He}) = 1$ ,

$p_3(\text{buy}|\text{He}, \text{can}) = 1$ ,  $p_3(\text{of}|\text{the}, \text{can}) = 1$ , ...,  $p_3(\cdot|\text{of}, \text{soda}) = 1$ .

- Entropy:  $H(p_1) = 2.75$ ,  $H(p_2) = .25$ ,  $H(p_3) = 0 \leftarrow \text{Great?!}$



# LM: an Example (The Problem)

- Cross-entropy:
- $S = \langle s \rangle \langle s \rangle$  It was the greatest buy of all.
- Even  $H_S(p_1)$  fails ( $= H_S(p_2) = H_S(p_3) = \infty$ ), because:
  - all unigrams but  $p_1(\text{the})$ ,  $p_1(\text{buy})$ ,  $p_1(\text{of})$  and  $p_1(\cdot)$  are 0.
  - all bigram probabilities are 0.
  - all trigram probabilities are 0.
- We want: to make all (theoretically possible\*) probabilities non-zero.

\*in fact, all: remember our graph from day 1?